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Analysis of the Characteristics of Monthly Rainfall Pattern in Katsina

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Authors' contributions

This work was carried out in collaboration among all authors. Author EVT designed the study, performed the statistical, spectral and nonlinear analysis of the data, wrote the protocol, and wrote the first draft of the manuscript. Authors FOA and EJ managed the analyses of the study. Author EJ managed the literature searches. All authors read and approved the final manuscript.

Article Information

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ABSTRACT

Aims: This paper seeks to analyse the characteristics of monthly rainfall pattern in Katsina City in a view to unveiling the trends and describing its dynamics so that adequate recommendations can be made for its modelling.

Study Design: The analysis involves a complete statistical, trend, spectral and nonlinear analysis of the monthly rainfall time series recorded in Katsina.

Place and Duration of Study: Location: Katsina City, Katsina State, Nigeria from 1990 to 2015; a period of 26 years.

Methodology: Secondary data of daily rainfall recorded in Katsina city from 1990 to 2015 was collected from the Nigerian Meteorological Agency (NiMet), and monthly averages were taken to obtain the monthly rainfall data. The data was then subjected to statistical, trend, spectral and nonlinear analysis techniques to reveal the behavioural patterns in the rainfall and also to reveal its underlying dynamics for its future modelling and prediction.

Results: The outcome of this analysis indicates that the monthly rainfall in Katsina exhibits an increasing trend with high variance and right-skewed distribution requiring a maximum of 6

independent variables to model its dynamics. The correlation exponent plot reached a saturation value of 5.892 confirming that the monthly rainfall in Katsina over the last 26 years exhibits low dimensional chaotic behavior while the largest Lyapunov exponent for the monthly rainfall time series in Katsina was also computed and found to be positive, having a value of 0.006055/month confirming the presence of deterministic chaos dynamics and is predictable for the next 165 months.

Conclusion: Since from the findings of this work it is confirmed that the rainfall in Katsina exhibits chaotic behavior with an increasing trend, it is recommended that more drainages and dams be built to provide steady supply of water for agricultural and domestic purposes as well as curtail the menace of flooding and drought which may occur as a result of global warming and climate change.

Keywords: Nonlinear dynamics; trend analysis; phase space reconstruction; phase portrait; correlation dimension; Lyapunov exponent.

1. INTRODUCTION

In recent times, scientists have laid more emphasis on the modelling of time series as a tool to ease the management and forecasting of the earth's meteorological and hydrological resources. Time series represents a dynamic measure of a physical process over a given period of time and may be discrete or continuous [1]. The discovery of Chaos by Edward Lorenz in 1961 [2], has brought about a great revolution in the mode of understanding and expressing most of these phenomena in nature. Chaos theory, the basis and foundation of nonlinear dynamics, is a tool that can be used for characterizing and modelling complex phenomena in nature such as rainfall data which has a higher variation coefficient [3]. Weather is a continuous, dataintensive, multidimensional, dynamic and chaotic process and these properties make weather prediction a big challenge as the chaotic nature of the atmosphere implies the need for massive computational power required to solve the equations that describe the atmospheric conditions [4]. Climate indeed varies nonlinearly too, but this has not prevented scientists from making good predictions using advanced regression techniques. Science and technology have been applied to predict the state of the atmosphere in future time for a given location, and this is very important as it affects life on earth. Today, computational weather forecasts are made by collecting quantitative data about the current state of the atmosphere and using scientific understanding of atmospheric processes to numerically project how the atmosphere will evolve, but due to an incomplete understanding of the chaotic atmospheric processes, forecasts become less accurate as the range of forecast increases [5].

This paper is focused on undertaking a detailed behavioural analysis of the monthly rainfall in Katsina over the last twenty-six years so as to unveil its dynamics thereby characterizing the data for modelling and forecasting to boost the planning of agricultural activities in the nearest future.

2. MATERIALS AND METHODS

The behavioral analysis of daily rainfall in Katsina state will be undertaken in this research by applying the following techniques: statistical analysis of the data, trend and spectral analysis, and nonlinear analysis.

2.1 Statistical Analysis

Statistical analysis involves the computation of the arithmetic mean, variance and standard deviation, the coefficient of variation, signal-tonoise ratio, range, kurtosis and skewness. Skewness is a measure of the asymmetry of the data around the sample mean. If skewness is negative, the data are spread out more to the left of the mean than to the right. If skewness is positive, the data are spread out more to the right. The skewness of the normal distribution (or any perfectly symmetric distribution) is zero. The skewness, S of a distribution with mean μ and standard deviation σ is given as [6]:

$$S = \frac{E(x-\mu)^3}{\sigma^3} \tag{1}$$

The parameter E(t) represents the expected value of the quantity t. Kurtosis, on the other hand, is a measure of how outlier-prone (scattered and detached) a distribution is. The kurtosis of the normal distribution is 3 while distributions that are more outlier-prone than the

normal distribution have kurtosis greater than 3; with distributions that are less outlier-prone have kurtosis less than 3. The kurtosis, K of a distribution with mean μ and standard deviation σ is given as [6]:

$$K = \frac{E(x-\mu)^4}{\sigma^4} \tag{2}$$

MATLAB statistics toolbox (R2014a) is used to achieve these computations.

2.2 Trend Analysis

In order to check the overall effect of the greenhouse effect and global warming on the rainfall pattern in Katsina, trend analysis was carried out using the following statistical tools:

- i. the correlation coefficient of the rainfall data with time was computed to determine the strength of the linear relationship the daily rainfall data with time,
- ii. the monotonic increasing or decreasing trend was tested using the non-parametric Mann-Kendall test, and
- iii. the slope of a linear trend is estimated with the nonparametric Sen's slope estimator.

2.2.1 Correlation coefficient

The Pearson product moment correlation coefficient R, is a parametric test that measures the strength and the pattern of a linear relationship between two variables. It is mathematically given by [7]:

$$R = \frac{n \sum xy - (\sum x)(\sum y)}{\sqrt{n(\sum x^2) - (\sum x)^2} \sqrt{n(\sum y^2) - (\sum y)^2}}$$
(3)

R value ranges from -1 to +1, with +1 or -1indicating a perfect correlation (positive or negative) and a correlation coefficient close to or equal to zero indicating no relationship between the variables. A correlation greater than 0.8 is generally described as strong, whereas a correlation less than 0.5 is generally described as weak. While a positive correlation coefficient indicates a positive relationship i.e. both variables move in same direction, a negative correlation coefficient indicates a negative relationship i.e. both variables move in opposite directions [8].

2.2.2 Mann-Kendall analysis

The nonparametric Mann-Kendall test is usually used to detect trends that are monotonic but not necessarily linear. The Mann-Kendall test statistic S is computed using the formula [9]:

$$S = \sum_{k=1}^{n-1} \sum_{j=k+1}^{n} sign(x_j - x_k),$$
 (4)

Where x_j and x_k are the daily rainfall values and time in days *j* and *k*, with j > k, respectively. The sign () function is defined as [10]:

$$sign(x_{j} - x_{k}) = \begin{bmatrix} 1 & if \ x_{j} - x_{k} > 0 \\ 0 & if \ x_{j} - x_{k} = 0 \\ -1 & if \ x_{i} - x_{k} < 0 \end{bmatrix}$$
(5)

A very high positive value of S (>120) is an indicator of an increasing trend, while a very low negative value indicates a decreasing trend [11]. The Man-Kendall parameter S and its variance VAR(S) are used to compute the test statistic Z as follows [9]:

$$Z = \begin{cases} \frac{S-1}{\sqrt{VAR(S)}} & if \ S > 0\\ 0 & if \ S = 0\\ \frac{S+1}{\sqrt{VAR(S)}} & if \ S < 0 \end{cases}$$
(6)

The Z statistic follows a normal distribution trend, is tested at 95% (α =0.05) level of significance ($Z_{\frac{\alpha}{2}} = 1.96$) and its value describes the trend as [12]:

- i. decreasing if Z is negative and the absolute value is greater than the level of significance,
- ii. increasing if Z is positive and greater than the level of significance, and
- iii. no trend if the absolute value of Z is less than the level of significance.

2.2.3 Sen's slope estimator

The Sen's test is estimates the linear regression coefficient or true slope of an existing trend (i.e. change per day). The Sen's method is used in cases where the trend can expressed as linear:

$$y(t) = Qt + B \tag{7}$$

Where Q is the slope, B is a constant and t is time.

Sen's estimator,
$$Q = \underbrace{median}_{j>k} \left(\frac{x_j - x_k}{j - k} \right)$$
 (8)

For *n* values x_j in the time series there will be as many as $N = \frac{n(n-1)}{2}$ slope estimates Q_i of which the median value gives the Sen's estimator, Q. In order to get an estimate of the intercept B in equation (7), the *n* values of differences $x_i - Qt_i$ are calculated and the median of these values gives an estimate of B [13].

2.3 Spectral Analysis

Spectral analysis is another simple way of characterizing attractors and is often used to qualitatively distinguish quasi-periodic or chaotic behavior from periodic structure and also to identify different periods embedded in a chaotic signal. Chaotic signals are characterized by the presence of wide broadband noise in their power spectrum, with a continuum of frequencies in their oscillations [14]. The power spectrum of a signal shows how a signal's power is distributed throughout the frequency domain [15]. To convert the rainfall time domain series to frequency domain, the fast Fourier transform (fft) was applied. The power per Hertz is obtained from the square of the absolute value of the fast Fourier transform [16]:

$$\frac{Power}{Hz} = abs\{fft[x(t)]\}^2$$
(9)

The periodicity of the rainfall in Katsina was estimated from the power spectrum as the reciprocal of the dominant frequency (peak or fundamental frequency) of the power spectrum plot [17].

2.4 Nonlinear Analysis

The tools of nonlinear analysis used to characterize the daily rainfall data in this paper include time series plot, phase portrait and Poincaré map, correlation dimension, Lyapunov exponents and Kolmogorov-Sinai entropy.

2.4.1 Time series plot

Time series plot involves plotting the daily, monthly and yearly rainfall data and observing the trend. If they exhibit irregular, aperiodic or unpredictable behaviour, then it could be described as random or chaotic. On the other hand, if they exhibit a regular repeating pattern, then the system exhibits either a periodic and quasi periodic behavior [18].

2.4.2 Phase portrait

A phase portrait is a two-dimensional visualization of the phase-space. It displays the attractor and unveils its dynamics. Chaotic systems exhibit distinct shapes, periodic systems exhibit limit cycle (closed curves) while quasi periodic systems exhibit torus shape [14].

2.4.3 Poincaré maps

The Poincaré map is that it represents a slice through the attractor of the dynamical system and it is a stroboscopic view of the phase portrait of the dynamical system; hence it can also be referred to as a stroboscopic map [19]. Poincaré maps of periodic systems show a single point, quasi-periodic systems show a closed curve while chaotic systems show distinct points. A summary of the different dynamical systems and their characteristics is shown in Table 1 [14].

Table 1. Different dynamic systems and the structure of their power spectrum, phase portraits and Poincaré maps

Solution of Dynamical System	Fixed	Periodic	Quasi Periodic	Chaotic
Power spectrum	-	Single dominant peak	Dominant peak and other sub- peaks	Broad band noise with continuum of frequencies; may peak at $f_0 = 0$
Phase portrait	Point	Closed Curve	Torus	Distinct Shapes
Poincaré Maps	-	Point	Closed Curve	Space filling or Ergodic points

2.4.4 Correlation dimension

The correlation dimension gives a measure of the complexity or number of active degrees of freedom excited by the system [20]. The Grassberger-Procaccia algorithm is used to compute the correlation dimension in this work using the correlation integral. For any set of M points in an *m*-dimensional phase space, the correlation integral or correlation sum (spatial correlation of points) C(r) is computed by the equation [21]:

$$C_m(r) = \lim_{N \to \infty} \frac{2}{N(N-1)} \sum_{i=1}^{M} \sum_{j=i+1}^{M} H(r - x_i - x_j)$$
(10)

H(x) is the Heaviside function and $\|...\|$ is the Euclidean norm, while r is the scaling parameter. The correlation integral measures the fraction of the total number of pairs of phase points that are within a distance r from each other. For chaotic time series, the correlation integral power law for small values of r takes the form:

$$C(r) \sim r^{\nu} \tag{11}$$

Thus, the correlation dimension v is given by:

$$\nu = \lim_{r \to 0} \lim_{M \to \infty} \frac{\log C(r)}{\log r}$$
(12)

Hence, a log-log graph of the correlation integral versus the scaling parameter, r will yield an estimate of the correlation dimension v, which is computed from the slope of a least-square fit of a straight line over a large length scale of r. For

chaotic systems, the correlation exponent curve for a range of values of embedding dimension (say m = 2 to 30) usually saturates at values beyond its actual embedding dimension. The saturation value of the correlation exponent plot gives the correlation dimension and the value of the embedding dimension at which the saturation of the correlation exponent curve occurs generally provides an upper bound on the number of variables sufficient to model the dynamics [18]. The dynamics of different systems based on their correlation exponent curve is described in Fig. 1 [3]:

In Fig. 1, it can be seen that for chaotic systems, the correlation exponent curve increases initially then saturates after a specific value of embedding dimension are reached, with the saturation value of the correlation exponent curve being the correlation dimension of the system. The correlation exponent plot of stochastic systems generates a monotonic increasing curve which never saturates to a specific value while deterministic systems generate a constant value which does not vary with increasing embedding dimension.

Furthermore, If the calculation of correlation dimension leads to a finite integer value, the underlying dynamics of the system is considered to be dominated by some strong periodic phenomena whereas if the value is fractional (and usually small) then the system is considered to be dominated by low dimensional deterministic chaotic dynamics governed by the geometrical and dynamical properties of the attractor [22].



Fig. 1. Characterization of systems based on their correlation exponent plot [3]

2.4.5 Lyapunov exponents

Lyapunov exponents (λ) are the average rates of exponential divergence or convergence of nearby orbits in phase space and is a fundamental property that characterizes the rate of separation of infinitesimally close trajectories [23]. It is mathematically given by:

$$\lambda_1(i) = \frac{1}{i\Delta t} \cdot \frac{1}{M-i} \cdot \sum_{j=1}^{M-i} \ln \frac{d_j(i)}{d_j(0)}$$
(13)

 Δt is the sampling period of the time series, M is the number of reconstructed phase points and $d_j(i)$ is the distance between the *j*th pair of nearest neighbors after *i* discrete-time steps, i.e., $i.\Delta t$ seconds. The nearest neighbor, X_j , is found by searching for the point that has the least distance to the particular reference point, X_j . This is expressed as:

$$d_{j}(0) = \min_{X_{j}} \|X_{j} - X_{j}\|$$
(14)

 $d_j(0)$ is the initial distance from the *jth* point to its nearest neighbor \hat{j} . A positive Lyapunov exponent indicates chaotic behavior, a negative value indicates a dissipative system i.e. a stable fixed point while a zero Lyapunov exponent indicates conservative system i.e. a periodic one or stable limit cycle [24]. The method used in this work to compute the largest Lyapunov exponent was developed by Rosenstein et al. in 1993 [25]. The Lyapunov (error folding) time or predictability T is the inverse of the largest Lyapunov exponent λ and is expressed as [26]:

$$T = \frac{1}{\lambda_1} \tag{15}$$

2.4.6 Phase space reconstitution

In order to effectively carry out nonlinear analysis, phase space reconstruction has to be done so as to draw out a multi-dimensional description of the system in an embedded space called state space. The method of delays was thus employed to achieve this [23],[27]. For a generalized time series $\{x_1, x_2,...,x_N\}$, the attractor can be reconstructed in a m-dimensional phase space of delay coordinates in form of the vectors:

$$X_n = [x_n, x_{n+\tau}, x_{n+2\tau}, \dots, x_{n+(m-1)\tau}]$$
(16)

 τ is the time lag, and m is the embedding dimension. The time delay τ is evaluated in this work using the method of average mutual

information (AMI) developed by Cellucci *et al.* in 2003 [28]. In order to obtain the time delay, the value of the lag length at the first local minimum of the AMI plot corresponding to the delay time of the time series [3],[18]. The minimum embedding dimension, m was computed using the method of "False Nearest Neighbors (FNN)" which was developed by Kennel *et al.* in 1992 [29]. By plotting the percentage of FNN against increasing embedding dimension values, a monotonic decreasing curve is observed and the minimum embedding dimension can be evaluated from the point where the percentage of FNN drops to almost zero or a minimum value.

The mean period, P of the data was computed as the inverse of the peak period of the fast Fourier transform. The mean period or periodicity P in a time series removes cyclic/seasonal variations in a time series data by seasonal differencing technique. Seasonal differencing is achieved by taking the difference between an observation and the corresponding observation from the previous year or season using the expression [30]:

$$x_t' = x_t - x_{t-S}$$
(17)

Where S = number of seasons or periodicity of the data. For monthly data, S = 12.

The phase space reconstruction will not be properly achieved and the deterministic components of the data will not be adequately revealed if the data is not made stationary and this could lead to misleading results in the nonlinear analysis of the data [31].

2.5 Study Area and Data Source

Katsina state, also known as the home of hospitality, is located in the North-Western region of Nigeria. The state is located within the coordinates 12°15'N, 7°30'E and 12°25' N, 7°50' E, and was created on 23rd September 1987. It covers a total land area of 24,192 km² with a population density of 160 /km² and its landscape is largely dominated by the Sahel savannah vegetation. Katsina state experiences two dominant seasons: the rainy and dry season, with the Hausa-Fulani who are predominantly farmers being the largest ethnic group in the state [32]. The data used in this research was obtained from the Nigerian Meteorological Agency (NiMet) Abuja. lt comprises of secondary data made up of daily average rainfall (mm) recorded in Katsina from 1st January 1990 to 31st December 2015, a

period of twenty-six years. The daily rainfall values were then converted to monthly rainfall by taking the mean value per month.

3. RESULTS AND DISCUSSION

The results of the behavioural analysis of rainfall pattern in Katsina is presented in this section.

3.1 Results of the Statistical Analysis

The statistics of daily rainfall (mm) is displayed in Table 2:

Table 2. Statistics of daily rainfall in Katsina

Statistic	Value
No. of data points	312
No. of nill values	153 (49%)
Mean (mm)	1.573
Standard Deviation (mm)	2.349
Variance (mm)	5.518
Coefficient of Variation (cv)	1.494
Signal-to-noise ratio	0.669
Maximum value (mm)	11.60
Minimum value (mm)	0.00
Kurtosis	5.253
Skewness	1.660

The results in Table 2 show a generally low overall mean value of daily rainfall (1.57 mm), a high variance of 5.518 and standard deviation of 2.349 (cv = 1.494). Furthermore a kurtosis of 5.253 and Skewness of 1.66 (skew to the right) with a significant number of null rainfall values (49%) in the data indicating a sparse irregular

months

(a)

12

10

ainfall (mm)-

distribution (high outlier-prone data) of monthly rainfall in Katsina over the last 26 years. This is attributed to the fact that Katsina is located in the Sahel savannah region of Nigeria within the Sahara desert region, hence the limited and sparse amount of rainfall received in the town. Fig. 2. (a) & (b) shows time series plots of monthly and seasonally differenced monthly rainfall in Katsina.

3.2 Trend Analysis

The summary of the trend analysis of the converted annual rainfall data using the Mann-Kendall trend test, Sen's slope estimator and Pearson's correlation coefficient are displayed in Table 3 and Fig. 3. The seasonal monthly rainfall values were used here by removing the null values recorded during the dry season.

The trend analysis results in Table 3 (Mann-Kendall test) and Fig. 2 (Sen's slope estimator) indicates that the trends of the annual rainfall in Katsina is significant as the Z-statistic computed (2.397) is greater than the z-value at the level of significance (1.96). This implies an increasing trend in the seasonal monthly rainfall in Katsina City. Hence there could be an increased risk of occurrences of flooding and surface run-off/erosion in the nearest future.

3.3 Results of Spectral Analysis

The result of the spectral analysis of daily rainfall in Katsina from 1990 to 2015 is displayed in Fig. 4.



Fig. 2. Time Series for: (a) Monthly rainfall and (b) Seasonally differenced monthly rainfall (seasonality = 12 months) time series for Katsina from 1990-2015

Variable	Result
Pearson's correlation coefficient (R)	0.2025
Kendall tau	0.1278
Mann-Kendell coefficient S	1625
Z statistic	2.397
Hypothesis test (h=1: significant, h=0: not significant)	h = 1
Trend description (from R and Z values)	Increasing trend
Trend Significance	Significant

Table 3. Summary of the Mann-Kendall analysis for annual rainfall in Katsina



Fig. 3. Seasonal monthly rainfall trend for Katsina using Sen's Slope Model, y = 0.084t + 1.866 (increasing trend observed)



Fig. 4. Power Spectrum of Rainfall in Katsina showing the Dominant Frequency

The result of the spectral analysis displayed in Fig. 4 shows that the rainfall in Katsina has a single dominant peak with a mean annual monthly cycle of 12 months rainfalls. Also the

average rainfall duration per year which is the inverse of the frequency of the second dominant peak in the power spectrum was found to be about $5.25 \approx 5$ months.

3.4 Results of the Nonlinear Analysis

Fig. 5 shows the estimation of time lag using the method of average mutual information (AMI). A delay time of 2 months was calculated for the monthly rainfall dataset. Fig. 6, on the other hand, illustrates the determination of the optimum embedding dimension using the method of false nearest neighbours (FNN). The monthly rainfall data for Katsina was found to have an embedding dimension of eleven (m = 6). The embedding dimension value obtained indicates that the rainfall in Katsina requires a maximum of 6 independent variables (degrees of freedom) to model its dynamics.

Figs. 7 and 8 show the phase portrait and Poincaré map for rainfall constructed using the time lag and embedding dimensions calculated. The phase portrait exhibits a sponge-like geometry of distinct shapes tending towards the origin (zero) while the Poincaré map shows scattered distinct points also tending towards an equilibrium point (attractor) indicating the presence of dissipative-damped random cycles in the dynamics of the rainfall time series. These plotted phase points are concentrated at the origin due to the number of null values (49%) in the monthly rainfall dataset which is as a result of the sparse distribution of rainfall in Katsina.

The correlation dimension was then calculated for the monthly rainfall data using the time lag $\tau = 2$ and for increasing embedding dimensions, *m*, from 2 to 45. Fig. 9 is a plot showing the relationship between the correlation function *C*(*r*) and the radius *r* (i.e. $\log C(r)$ versus $\log r$) for increasing embedding dimension *m* while Fig. 10 shows the relationship between the correlation exponents and the embedding dimension values *m*. It is observed from Fig. 10 that the correlation exponent values increase with increasing embedding dimension up to m = 15 and then plateaus and saturates to a value v = 5.892. The saturation of the correlation exponent indicates a likely chaotic behavior in the monthly rainfall time series with a correlation dimension of 5.89. This implies that six (6) degrees of freedom are required to model the monthly rainfall in Katsina city.

The Lyapunov spectrum obtained from the computation of the Lyapunov exponent for monthly rainfall in Katsina using Rosenstein's algorithm is displayed in Fig. 11 while the details of the Lyapunov exponents for increasing values of embedding dimension is presented in Table 4.

Table 4. The Lyapunov exponent values from m=1 to 6

Embedding	Lyapunov exponent		
dimension (m)	(λ)		
1	0.002319599722999		
2	0.002612172302405		
3	0.004597975112222		
4	0.005108843110237		
5	0.005288871379006		
6	0.006055059727202		
*1 (1			

*Largest Lyapunov exponent, $\lambda = 0.006055$ /month

The largest Lyapunov exponent for the monthly rainfall time series in Katsina was computed and found to be 0.006055/month. The positive value of the largest Lyapunov exponent confirms the fact that the monthly rainfall in Katsina over the last 26 years exhibits chaotic behaviour and the monthly rainfall is predictable for the next 165 months.



Fig. 5. Estimation of time lag using the method of AMI (T = 2 months)







Fig. 7. Phase portrait of monthly rainfall in Katsina



Fig. 8. Poincaré map of monthly rainfall in Katsina



Fig. 9. log-log plot showing the relationship between the Correlation Integral C(r) and the Scaling Radius r for different values of embedding dimension for monthly rainfall in Katsina from 1990-2015



Fig. 10. The relationship between correlation exponent and embedding dimension *m* for monthly rainfall in Katsina from 1990-2015



Fig. 11. The Lyapunov Spectrum for the Estimation of the Largest Lyapunov Exponent for monthly rainfall in Katsina from 1990-2015.

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4. CONCLUSION

In this paper, an analysis of the characteristics of monthly rainfall pattern in Katsina from the year 1990-2015 was carried out. The outcome of this analysis indicates that the rainfall in Katsina exhibits an increasing trend with high variance and low dimensional chaotic behaviour. A maximum of six (6) independent variables is required to model the monthly rainfall in Katsina while the rainfall is sparse and has good predictability in the next 165 months. It is recommended that adequate measures be taken to stem the trend of irregular rainfall pattern in the years to come. These measures include clearing of drainages and waterways as well as the building of dams to provide water for irrigation and to also help contain water from flash floods which is being envisaged from the results of this research as a result of climate change which is very eminent in the northern part of Nigeria.

COMPETING INTERESTS

Authors have declared that no competing interests exist.

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